

# Blind Source Separation Using Renyi's Mutual Information

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**Abstract**—A blind source separation algorithm is proposed that is based on minimizing Renyi's mutual information by means of nonparametric probability density function (PDF) estimation. The two-stage process consists of spatial whitening and a series of Givens rotations and produces a cost function consisting only of marginal entropies. This formulation avoids the problems of PDF inaccuracy due to truncation of series expansion and the estimation of joint PDFs in high-dimensional spaces given the typical paucity of data. Simulations illustrate the superior efficiency, in terms of data length, of the proposed method compared to fast independent component analysis (FastICA), Comon's minimum mutual information, and Bell and Sejnowski's Infomax.

**Index Terms**—Blind source separation, component analysis, Givens rotations, mutual information, Renyi's entropy, Renyi's information.

## I. INTRODUCTION

MINIMIZATION of the mutual information (MMI) between outputs is considered the ideal information theoretic criteria for blind source separation (BSS) [1]–[3]. The (Shannon) mutual information can be written as the sum of the (Shannon) marginal entropies minus the joint entropy. One of the difficulties of using MMI is the estimation of the marginal entropies. In order to estimate the marginal entropy, Comon and others approximate the output marginal probability density functions (PDFs) with truncated polynomial expansions [2]–[4], a process that inherently introduces error in the estimation procedure. Another MMI method proposed by Xu *et al.* [5] avoids the polynomial expansion by approximating the Kullback–Leibler divergence with the Cauchy–Schwarz inequality and estimates Renyi's entropy nonparametrically by Parzen windows. The second method requires estimation of the  $N$ -dimensional joint entropy, and nonparametric PDF estimation using Parzen windows is ill-posed in high-dimensional spaces [5]. We propose below a new algorithm based on Renyi's mutual information that requires only marginal entropy estimation and avoids both polynomial expansions and estimation of Renyi's joint entropy.

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## II. COST FUNCTION

Renyi's mutual information is defined as [6]

$$I_{R_\alpha}(y) = \frac{1}{\alpha - 1} \log \int_{-\infty}^{\infty} \frac{f_Y(y)^\alpha}{\prod_{i=1}^N f_{Y_i}(y_i)^{\alpha-1}} dy. \quad (1)$$

The sum of Renyi's marginal entropies minus the joint entropy is

$$\sum_{i=1}^N H_{R_\alpha}(y_i) - H_{R_\alpha}(y) = \frac{1}{\alpha - 1} \log \frac{\int_{-\infty}^{\infty} f_Y(y)^\alpha dy}{\int_{-\infty}^{\infty} \prod_{i=1}^N f_{Y_i}(y_i)^\alpha dy_i} \quad (2)$$

which differs from (1). However, both (1) and (2) are nonnegative [6], and both evaluate to zero when and only when the joint PDF can be written as a product of the marginals, which occurs identically when the statistically independent sources are separated. Therefore, minimizing Renyi's MI can be accomplished by, minimizing the sum of (Renyi's) marginal entropies minus the joint entropy, just as with Shannon's entropy. Notice that (2) is not a reformulation of MI, but is preferred since Renyi's entropy can be estimated nonparametrically as proposed in [9].

In order to produce an algorithm that scales favorably as the number of sources increases, the estimation of the joint entropy is avoided by the two-step parameterization proposed in [4]. The first stage performs spatial whitening, and the second stage performs a rotation in  $N$ -dimensions. The rotation, denoted as matrix  $R$ , is adapted to minimize the cost function given by (2). In fact, since Renyi's joint entropy is invariant to rotations

$$\begin{aligned} H_{R_\alpha}(y) &= \frac{1}{1 - \alpha} \log \int_{-\infty}^{\infty} f_Y(y) dy \\ &= \frac{1}{1 - \alpha} \log \int_{-\infty}^{\infty} \frac{f_X(x)^\alpha |\det R|}{|\det R|^\alpha} dx = H_{R_\alpha}(x) \end{aligned} \quad (3)$$

the joint entropy may be discarded from the adaptation process, leaving only the marginal entropies. Hence, the cost function becomes simply

$$J = \sum_{i=1}^N H_{R_\alpha}(y_i) \quad (4)$$

which mimics the cost function of [3], [4], but with Renyi's entropy substituted for Shannon's entropy. Using the nonparametric Renyi's entropy estimator proposed in [9], we arrive at a new algorithm for BSS.



Fig. 1. System block diagram for two sources/observations.

### III. SYSTEM DESCRIPTION

The overall BSS block diagram for two inputs/observations is given in Fig. 1. The equations for a system having  $N$  sources/observations are given by  $\mathbf{x} = \mathbf{W}^T \mathbf{H}^T \mathbf{s}$  and  $\mathbf{y} = \mathbf{R}^T \mathbf{x}$ , where  $\mathbf{W} = \Phi \Lambda^{-1/2}$ ,  $\Phi$  is the matrix of eigenvectors of the autocorrelation of  $\mathbf{H}^T \mathbf{s}$ , and  $\Lambda$  is the corresponding eigenvalue matrix. Notice that  $E[\mathbf{x}\mathbf{x}^T] = \mathbf{I}_N$  is the  $(N \times N)$  identity matrix, due to the spatial whitening. The rotation matrix  $\mathbf{R}$  is constructed from the product of  $N(N-1)/2$  Givens rotation matrices<sup>1</sup> are  $\mathbf{R}_{ij}$ , where  $\mathbf{R}_{ij}$  equals  $\mathbf{I}_N$  with elements  $\mathbf{I}_N(i, i)$ ,  $\mathbf{I}_N(i, j)$ ,  $\mathbf{I}_N(j, i)$ , and  $\mathbf{I}_N(j, j)$  modified to  $\cos \theta_{ij}$ ,  $\sin \theta_{ij}$ ,  $-\sin \theta_{ij}$ , and  $\cos \theta_{ij}$ , respectively, where  $\mathbf{I}_N(i, j)$  is the element of  $\mathbf{I}_N$  located at the  $i$ th row and  $j$ th column. There is one rotation angle for each  $\mathbf{R}_{ij}$ , whose purpose is to perform a rotation in the output space in the plane specified by the  $i$ th and  $j$ th (orthogonal) basis vectors. The gradient of  $\mathbf{R}$  with respect to  $\theta_{ij}$  is also needed for the algorithm development and it is denoted as  $\nabla \mathbf{R}_{ij}$ .

### IV. STOCHASTIC GRADIENT DESCENT ALGORITHM

When Parzen windowing is used with a Gaussian kernel, the estimate for Renyi's quadratic marginal entropy simplifies to [9]

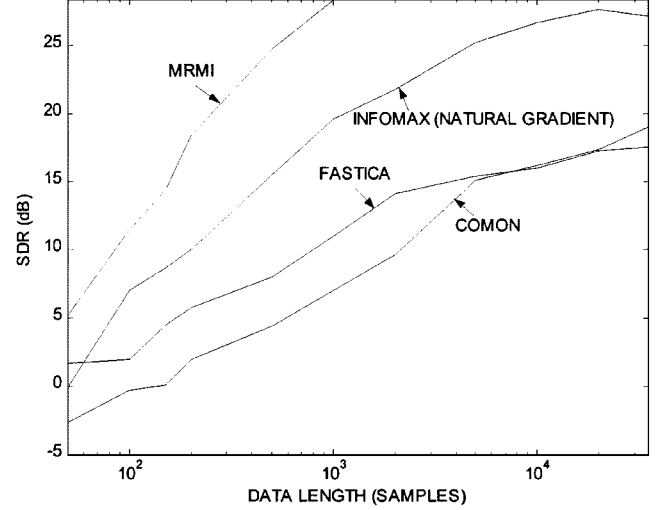
$$H_{R_2}(y_i) = -\log \frac{1}{L^2} \sum_{j=1}^L \sum_{k=1}^L G(y_i(j) - y_i(k), 2\sigma^2) \quad (5)$$

where  $G(x, 2\sigma^2)$  is a Gaussian PDF, and  $y_i(j)$  is the  $j$ th sample of output  $y_i$ . Notice that the infinite-limit integral disappears and that there are no approximations involved (aside from the implicit PDF estimation using Parzen windows). Substituting (5) into (4) and taking the derivative with respect to  $\theta_{ij}$  produces (6), shown at the bottom of the page, where  $(\nabla_{\theta_{ij}} \mathbf{R})_i$  is the  $i$ th column of  $\nabla_{\theta_{ij}} \mathbf{R}$ , and  $\mathbf{x}(m)$  is the vector of  $\mathbf{x}$  at time  $m$ . The overall update equation for stochastic gradient descent is then  $\theta(n+1) = \theta(n) - \eta \Delta \theta(n)$ , where  $\theta(n)$  and  $\Delta \theta(n)$  are vectors of angles, and  $\eta$  is the step size.

### V. RESULTS

The proposed method is compared to FastICA [7], Bell and Sejnowski's Infomax [8], and Comon's MMI method [4] using

<sup>1</sup>This is the number of unique pairwise combinations of  $N$  orthogonal basis vectors.

Fig. 2. Signal-to-distortion plots for ten sources/observations. Notice that the  $x$ -axis is logarithmic.

an instantaneous mixture of ten audio sources consisting of one music source and speech from five male and four female speakers. Spatial prewhitening is used for each method, the values of the mixing coefficients are chosen uniformly in  $[-1, 1]$  and the performance criterion is the signal-to-distortion ratio, defined as

$$\text{SDR} = \frac{1}{N} \sum_{i=1}^N 10 \log_{10} \left( \frac{(\max q_i)^2}{q_i q_i^T - (\max q_i)^2} \right) \quad (7)$$

where  $\mathbf{q} = \mathbf{R}^T \mathbf{W}^T \mathbf{H}^T$ ,  $\mathbf{q}_i$  is the  $i$ th row of  $\mathbf{q}$ , and  $\max(\mathbf{q}_i)$  equals the maximum of the argument. This criterion effectively measures the distance of  $\mathbf{q}$  from the product of a permutation matrix with a diagonal matrix.

Fig. 2 shows the SDR for each method as a function of  $L$ , the data length in number of samples. FastICA uses the symmetric approach and the cubic nonlinearity, Infomax uses Amari's natural gradient [11], Comon's method uses a fourth-order PDF expansion (requiring estimates of third and fourth-order cumulants) and our algorithm, MRMI, uses a kernel size  $\sigma^2$  of 0.5 ( $L \leq 100$ ) or 0.25 ( $L > 100$ ). In addition, small step sizes were used to maximize SDR. As can be seen from the figure, the MRMI method requires much less data to yield a given performance level (good separation is achieved at an SDR of 20 dB). This is imperative in a number of applications such as channel equalization and in nonstationary environments. In hearing aid applications for example, Torkkola explains [10] that "Any system which attempts to ... apply some means of inverse filtering would have to be adaptable on almost a frame-by-frame basis to be effective." The results for MRMI are not shown beyond  $L = 1000$  due to the computational

$$\Delta \theta_{ij} = \frac{\partial}{\partial \theta_{ij}} \sum_{k=1}^N H_{R_2}(y_k) = - \frac{\sum_{m=1}^N \sum_{n=1}^L \sum_{k=1}^L G(y_k(m) - y_k(n), 2\sigma^2) (y_k(m) - y_k(n)) (\nabla_{\theta_{ij}} \mathbf{R})_i^T (x(m) - x(n))}{\sum_{m=1}^L \sum_{n=1}^L G(y_k(m) - y_k(n), 2\sigma^2)} \quad (6)$$

complexity, which is  $O(L^2)$  compared to  $O(L)$  for the other three methods.

## VI. CONCLUSIONS

A new BSS algorithm has been developed that is based on the minimization of Renyi's mutual information. Unlike the Comon approach, which uses Shannon's entropy and requires truncation of a PDF series expansion, there are no approximations in the proposed method due to the utilization of Renyi's quadratic entropy. While this algorithm is computationally more complex than its predecessors, it has been shown to be much more efficient in terms of the amount of data samples required, an item that is paramount in tracking a rapidly changing environment.

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# Correspondence

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## Corrections to “Blind Source Separation Using Renyi’s Mutual Information”

In the above paper [1], the authors have found that it contains an incorrect statement. The introductory section includes the sentence “However, both (1) and (2) are nonnegative [6], and both evaluate to zero when and only when the joint PDF can be written as a product of the marginals.” As it is written, reference [6] in the letter appears to apply to both (1) and (2); however, it should only apply to (1). Furthermore, the previously held belief that the estimator of Renyi’s Mutual Information, given by (2), can never be negative has since been determined to be in error. This is the case, for example, for sub-Gaussian sources. Although it has been proven that (2) has a minimum value of zero for the case that there are two Laplacian sources, there is no known proof that generalizes this result for any combination of super-Gaussian sources and/or for any number of sources. It does appear that, for sub-Gaussian sources, there is a local minimum at the location corresponding to separation (which evaluates to zero), but it may not be the global minimum. Therefore, the sentence above should be corrected to read “... (1) is nonnegative and evaluates to zero when and only when the joint PDF can be written as a product of the marginals [6]. In addition, for super-Gaussian sources there is experimental evidence that (2) is also nonnegative and evaluates to zero only when the joint PDF can be written as a product of the marginals.” The end result is that the criterion must be slightly modified in order for it to also separate sub-Gaussian sources. The details of the modification may be found in a paper by the same authors [2]. The authors wish to extend an apology for any inconvenience this may have caused.

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## Comments on “Fast Approximation of Kullback–Leibler Distance for Dependence Trees and Hidden Markov Models”

For the above paper [1], the Associate Editor coordinating the review of this manuscript and approving it for publication was Dr. Marcelo A. Bruno.

### REFERENCES

- [1] M. N. Do, “Fast approximation of Kullback–Leibler distance for dependence trees and Hidden Markov models,” *IEEE Signal Processing Lett.*, vol. 10, pp. 115–118, Apr. 2003.

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## Corrections to “Robust Stability of Two-Dimensional Uncertain Discrete Systems”

The authors of the above paper [1] point out that, unfortunately, they were unaware that a similar topic has been studied in [2], and that Theorem 1 and Theorem 2 of the letter have already been obtained in [2].

H. Kar and V. Singh (Dept. Electrical-Electr. Engineering, Atilim Univ., Incek, Ankara, Turkey) also made a comment on the following statement appearing on page 135 of the letter [1]: “Give an initial value  $\alpha > 0$ ,  $[\cdot \cdot \cdot]$  small enough. Solve LMI (20) again. If  $\alpha$  approaches one eventually, then it means that the uncertain discrete-time 2-D system is *not* robustly asymptotically stable.”

Since the Lyapunov theorem provides only a sufficient condition for stability, the last sentence in this statement is not justified. In other words, merely on the basis of the fact that the conditions given in Theorem 2 are not met, it would not be correct to draw the inference that the uncertain discrete 2-D system is not robustly asymptotically stable.

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